

CIRJE-F-652

The Ten Commandments for Optimizing Value-at-Risk and Daily Capital Charges

Michael McAleer
Erasmus University Rotterdam
and Tinbergen Institute
and CIRJE, Faculty of Economics, University of Tokyo

August 2009

CIRJE Discussion Papers can be downloaded without charge from:

<http://www.e.u-tokyo.ac.jp/cirje/research/03research02dp.html>

Discussion Papers are a series of manuscripts in their draft form. They are not intended for circulation or distribution except as indicated by the author. For that reason Discussion Papers may not be reproduced or distributed without the written consent of the author.

The Ten Commandments for Optimizing Value-at-Risk and Daily Capital Charges*

Michael McAleer

Econometric Institute
Erasmus School of Economics
Erasmus University Rotterdam
The Netherlands

and

Tinbergen Institute
The Netherlands

and

Center for International Research on the Japanese Economy (CIRJE)
Faculty of Economics
University of Tokyo

Revised: August 2009

* The author wishes to acknowledge detailed comments from the Editor and a referee, and helpful discussions with Dave Allen, Manabu Asai, Federico Bandi, Peter Boswijk, Massimiliano Caporin, Felix Chan, Chialin Chang, Philip Hans Franses, Kees Jan van Garderen, Abdul Hakim, Shawkat Hammoudeh, Peter Reinhard Hansen, Eric Hillebrand, Juan-Ángel Jiménez-Martin, Masahito Kobayashi, Offer Lieberman, Shiqing Ling, Essie Maasoumi, Colin McKenzie, Marcelo Medeiros, Les Oxley, Teodosio Pérez-Amaral, Hashem Pesaran, Bernardo da Veiga and Alan Wong, and the financial support of the Australian Research Council. Previous versions of this paper have been presented at the Bank of Thailand, Catholic University of Brasilia, Chiang Mai University, Complutense University of Madrid, Erasmus University Rotterdam, Feng Chia University, Taiwan, Hong Kong Baptist University, Hong Kong University of Science and Technology, Keio University, Kyoto University, LaTrobe University, National Chung Hsing University, Taiwan, Pontifical Catholic University of Rio de Janeiro, Soka University, Japan, Stock Exchange of Thailand, Yokohama National University, the Universities of Amsterdam, Balearic Islands, Canterbury, Groningen, Padova, Palermo, Tokyo, Venice, Vigo and Zaragoza, and conferences in Chiang Mai, Thailand, Canterbury, New Zealand, Yokohama, Japan, Xiamen, China, and Taichung, Taiwan.

Abstract

Credit risk is the most important type of risk in terms of monetary value. Another key risk measure is market risk, which is concerned with stocks and bonds, and related financial derivatives, as well as exchange rates and interest rates. This paper is concerned with market risk management and monitoring under the Basel II Accord, and presents Ten Commandments for optimizing Value-at-Risk (VaR) and daily capital charges, based on choosing wisely from: (1) conditional, stochastic and realized volatility; (2) symmetry, asymmetry and leverage; (3) dynamic correlations and dynamic covariances; (4) single index and portfolio models; (5) parametric, semiparametric and nonparametric models; (6) estimation, simulation and calibration of parameters; (7) assumptions, regularity conditions and statistical properties; (8) accuracy in calculating moments and forecasts; (9) optimizing threshold violations and economic benefits; and (10) optimizing private and public benefits of risk management. For practical purposes, it is found that the Basel II Accord would seem to encourage excessive risk taking at the expense of providing accurate measures and forecasts of risk and VaR.

Keywords and phrases: Daily capital charges; excessive risk taking; market risk; risk management; value-at-risk; violations.

JEL Classifications: G32, G11, G17, C53.

“We regulators are often perceived as constraining excessive risk-taking more effectively than is demonstrably possible in practice. Except where market discipline is undermined by moral hazard, owing, for example, to federal guarantees of private debt, private regulation generally is far better at constraining excessive risk-taking than is government regulation.”

Alan Greenspan, Conference on Bank Structure and Competition, 8 May 2003

Elizabeth Turner: “Wait! You have to take me to shore. According to the Code of the Order of the Brethren ...”

Captain Barbossa: “First, your return to shore was not part of our negotiations nor our agreement so I must do nothing. And secondly, you must be a pirate for the pirate’s code to apply and you’re not. And thirdly, the code is more what you’d call “guidelines” than actual rules.”

Pirates of the Caribbean: Curse of the Black Pearl

1. Introduction

The cataclysmic financial meltdown worldwide that seems to have started in September 2008 has made it manifestly obvious that self-regulation in the finance and banking industry has been far from adequate. A period of soul searching is likely to be followed by much needed regulatory changes in the industry. Some changes to regulations governing the finance industry, especially as regards the monitoring and management of excessive risk, overseeing new financial instruments, and increased regulation of banks, are likely to be warranted, whereas others may have little or no effect.

Following the 1995 amendment to the Basel Accord (see Basel Committee on Banking Supervision (1988, 1995, 1996)), banks were permitted to use internal

models to calculate their Value-at-Risk (VaR) thresholds (see Jorion (2000) for a detailed discussion of VaR). This amendment was in response to widespread criticism that the standard approach to calculate and forecast VaR thresholds led to excessively conservative forecasts and higher mean daily capital charges that are associated with higher perceived risk. In short, the Basel II Accord was intended to encourage risk taking through self regulation.

Despite the well known antipathy of Alan Greenspan, the former Federal Reserve Chairman, to government regulation of excessive risk taking, it is no longer possible to argue logically that self regulation alone is a viable option in the banking and finance industry.

The primary purpose of this paper is to evaluate how best to forecast VaR and to optimize daily capital charges in an attempt to manage excessive risk taking as efficiently as possible, and to offer some idiosyncratic suggestions and guidelines regarding how improvements might be made regarding practical strategies for risk monitoring and management, especially when there is a strong and understandable movement into holding and managing cash (circa late-2008) than in dealing with risky financial investments.

The plan of the remainder of the paper is as follows. Section 2 presents the optimization problem facing authorized deposit-taking institutions. The data for a brief empirical analysis of VaR and daily capital charges are illustrated in Section 3. Section 4 provides a brief discussion of regression models and volatility models. Ten reasons for modelling time-varying variances, covariances and correlations using high and ultra high frequency data are given in Section 5. The Ten Commandments for

optimizing VaR and daily capital charges are discussed in Section 6. Some concluding comments are given in Section 7.

2. The Optimization Problem for Authorized Deposit-taking Institutions

Value-at-Risk (VaR) may be defined as “a worst case scenario on a typical day”. As such, it is concerned with relatively unlikely, or extreme, events. Given the financial turmoil in 2008, especially after September 2008, extreme events have become more commonplace, such that an extreme event would probably now need to be catastrophic to qualify as such.

Insofar as financial meltdowns tend to encourage Authorized Deposit-taking Institutions (ADIs) to shift financial assets into cash, VaR should still be optimized. However, ADIs may decide not to entertain the excessive risk associated with financial assets by holding a higher proportion of their portfolio in cash and/or relatively low risk assets.

Under the Basel II Accord, VaR forecasts need to be provided to the appropriate regulatory authority (typically, a central bank) at the beginning of the day, and is then compared with the actual returns at the end of the day.

For purposes of the Basel II Accord penalty structure for violations arising from excessive risk taking, a violation is penalized according to its cumulative frequency of occurrence in 250 working days, and is given in Table 1.

A violation is defined as follows:

Definition: A violation occurs when $VaR_t >$ negative returns at time t .

As encouraged by the Basel II Accord, the optimization problem facing ADIs, with the number of violations and forecasts of risk as endogenous choice variables, is as follows:

$$\text{Minimize } DCC_t = \sup_{\{k, VaR_t\}} \left\{ -(3+k)\overline{VaR}_{60}, -VaR_{t-1} \right\}, \quad (1)$$

where

DCC = daily capital charges,

VaR_t = Value-at-Risk for day t ,

$$VaR_t = \hat{Y}_t - z_t \cdot \hat{\sigma}_t,$$

\overline{VaR}_{60} = mean VaR over the previous 60 working days,

\hat{Y}_t = estimated return at time t ,

z_t = 1% critical value at time t ,

$\hat{\sigma}_t$ = estimated volatility at time t ,

$0 \leq k \leq 1$ is a violation penalty (see Table 1).

As \hat{Y}_t is typically difficult to predict, and as z_t is unlikely to change significantly especially on a daily basis), changes in $\hat{\sigma}_t$ are crucial for modelling VaR. Substantial

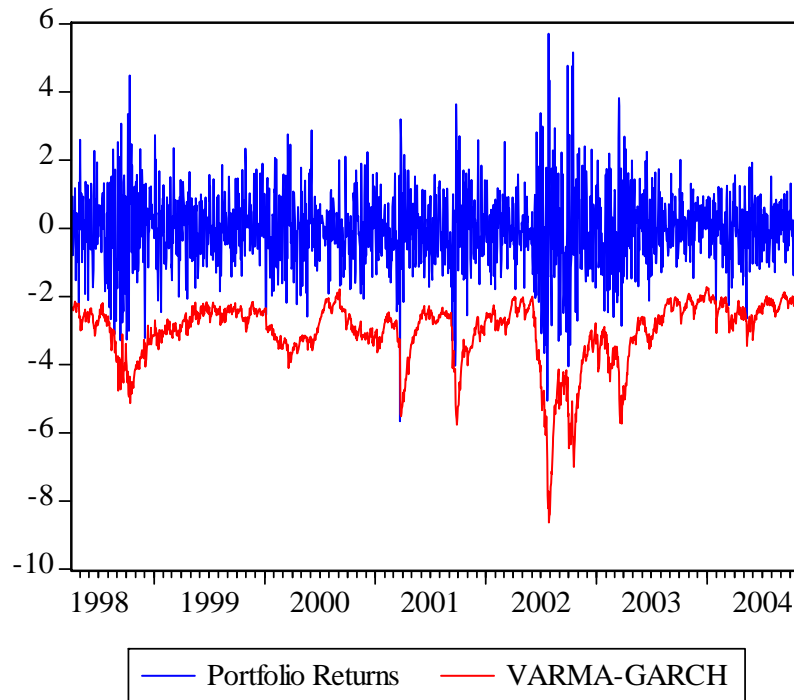
Table 1: Basel Accord Penalty Zones

Zone	Number of Violations	k
Green	0 to 4	0.00
Yellow	5	0.40
	6	0.50
	7	0.65
	8	0.75
	9	0.85
Red	10+	1.00

Note: The number of violations is given for 250 business days.

Note: The penalty structure under the Basel II Accord is specified for the number of penalties and not their magnitude, either individually or cumulatively.

Figure 10: Realized Returns and VARMA-GARCH VaR Forecasts



Note: This is Figure 10 in McAleer and da Veiga (2008a).

research has been undertaken in recent years to develop univariate and multivariate models of volatility, under the conditional, stochastic and realized volatility frameworks, in order to estimate $\hat{\sigma}_t$. Although VaR may also be estimated directly using regression quantiles (see, for example, Engle and Manganelli (2002)), this approach is not as popular as modelling volatility and then calculating VaR.

Although considerable research has been undertaken on higher-order moments, especially in the context of conditional volatility models, these are not considered in the paper. Model uncertainty for univariate and multivariate processes is also not considered (see Pesaran, Schleicher and Zaffaroni (2008) for an analysis of model averaging techniques for portfolio management).

As discussed above, the amendment to the Basel Accord was designed to reward institutions with superior risk management systems. For testing performance, a backtesting procedure, whereby the realized returns are compared with the VaR forecasts, was introduced to assess the quality of a bank's internal model. In cases where the internal models lead to a greater number of violations than could reasonably be expected, given the confidence level, the bank is required to hold a higher level of capital (see Table 1 for the penalties imposed under the Basel Accord).

If a bank's VaR forecasts are violated more than nine times in any financial year (see Table 1), the bank may be required to adopt the standard approach. As discussed in McAleer and da Veiga (2008b), the imposition of such a penalty is severe as it affects the profitability of the bank directly through higher daily capital charges, has a damaging effect on the bank's reputation, and may lead to the imposition of a more

stringent external model to forecast VaR thresholds, which would have the result of increasing daily capital charges for ADIs.

It should be noted that DCC is to be minimized, with k and VAR_t as endogenous choice variables. [The acronym DCC should be distinguished from a widely used multivariate conditional volatility model, which will be discussed below.]

It is worth giving a caveat with respect to the minimization of daily capital charges in times of extreme financial fluctuations, such as the period starting in September 2008. When excessive risk is very high, and is changing dramatically on a daily basis, daily capital charges should still be minimized, but sensible portfolio management may involve greater access to cash after VaR and daily capital charges have been determined.

3. Data

The data used in the empirical applications in McAleer and da Veiga (2008a, 2008b), and given as portfolio returns in Figure 10 above, are daily prices measured at 16:00 Greenwich Mean Time (GMT) for four international stock market indexes, namely S&P500 (USA), FTSE100 (UK), CAC40 (France), and SMI (Switzerland). All prices are expressed in US dollars. The data were obtained from DataStream for the period 3 August 1990 to 5 November 2004. At the time the data were collected, this period was the longest for which data on all four variables were available.

The synchronous returns for each market i at time t (R_{it}) are defined as:

$$R_{it} = \log(P_{i,t} / P_{i,t-1}),$$

where $P_{i,t}$ is the price in market i at time t , as recorded at 16:00 GMT.

For the empirical analysis, it was assumed that the portfolio weights are equal and constant over time, although both of these assumptions can easily be relaxed. The various conditional volatility models were used to estimate the variance of the portfolio directly for the single index model, and to estimate the conditional variances and correlations of all assets and asset pairs to calculate the variance of the portfolio for the portfolio model. Apart from the Standardized Normal and RiskmetricsTM models, all the conditional volatility models were estimated under the assumption that the distribution of the unconditional shocks was (1) normal; and (2) t , with 10 degrees of freedom.

The empirical results will be discussed further in the context of the last two Commandments, but the following points should be highlighted:

- (i) There is a trade-off between the number of violations and daily capital charges, with a higher number of violations leading to a higher penalty and lower daily capital charges through lower VaR.
- (ii) Apart from Standardized Normal, which does not estimate any parameters, and RiskmetricsTM, where the parameters are calibrated rather than estimated, the number of violations is higher for single index models than for their portfolio model counterparts, and the mean daily capital charges are correspondingly higher.

(iii) The use of the t distribution with 10 degrees of freedom always leads to fewer violations, and hence higher mean daily capital charges, for both the single index and portfolio models.

The two curves in Figure 10, as reproduced below, represent daily returns on the portfolio of the four aggregate indexes, as well as the forecast VaR thresholds using the VARMA-GARCH model of Ling and McAleer (2003). From the figure, it can be seen clearly that there are very few violations, which could mean one or more of the following:

- (1) the volatility model provides an accurate measure of risk and VaR;
- (2) the volatility model provides a conservative measure of risk and VaR;
- (3) the number of violations is not being used as an endogenous choice variable.

4. Regression Models and Volatility Models

The purpose of a (linear) regression model is to explain the conditional mean, or first moment, of Y_t :

$$Y_t = \alpha + \beta X_t + \varepsilon_t, \quad (2)$$

such that

$$E(Y_t|X_t) = \alpha + \beta X_t,$$

where α and β are taken to be scalar parameters for convenience.

The purpose of a (univariate) volatility model is to explain the variance, or second moment, of Y_t in (2):

$$\varepsilon_t = \eta_t \sqrt{h_t}, \quad \eta_t \sim iid(0,1) \quad (3)$$

where ε_t is the unconditional shock to the variable of interest, Y_t (typically, a stock return in empirical finance), which thereby has a risk component, η_t is the standardized residual (namely, a riskless shock, or fundamental), and $\sqrt{h_t}$ denotes volatility (or risk). One of the primary purposes in modeling volatility is to determine h_t to enable η_t to be calculated from the observable returns shock, ε_t . The volatility, h_t , may be latent or observed (though it is also likely to be subject to measurement error).

In the class of conditional volatility models, which are widely used for analysing monthly, weekly and daily data (though they can also be used for hourly data), h_t is modelled as:

$$h_t = \omega + \alpha \varepsilon_{t-1}^2 + \beta h_{t-1}, \quad (4)$$

where there are (typically sufficient) conditions on the parameters ω, α, β , to ensure that conditional volatility, h_t , is positive. The specification for h_t in (4) is the widely used generalized autoregressive conditional heteroskedasticity, or GARCH(1,1), model of Engle (1986) and Bollerslev (1986). As the model is

conditional on the information set at time $t-1$, current shocks do not affect h_t . For recent surveys of univariate and multivariate conditional volatility models, see Li, Ling and McAleer (2002), McAleer (2005a), and Bauwens, Laurent and Rombouts (2006).

Stochastic volatility models can incorporate leverage directly through the negative correlation between the returns and subsequent volatility shocks (see, in particular, Zhang, Mykland and Aït-Sahalia (2005), Aït-Sahalia, Mykland and Zhang (2006), and Barndorff-Nielsen, Hansen, Lunde Shephard (2008)). For a recent review of a wide range of stochastic volatility models, see Asai, McAleer and Yu (2006).

Realized volatility models can incorporate leverage as this is not explicitly excluded in the specification. However, realized volatility models are typically not specified to incorporate leverage.

At the multivariate level, it is not entirely clear how to define leverage in the context of conditional, stochastic or realized volatility models (for further details, see McAleer and Medeiros (2008a)).

5. Ten Reasons for Modelling Time-varying Variances, Covariances and Correlations Using High and Ultra High Frequency Data

Regression models are used to explain the conditional first moment of Y_t when it is not constant. Similarly, volatility models are used to explain the second moment of Y_t when it is not constant. It is well known that neither the first nor second moments of Y_t is constant, especially in the case of financial returns data.

Ten reasons for modelling time-varying variances, covariances and correlations using high frequency (namely, monthly, weekly, daily and hourly) and ultra high frequency (namely, minute and second) data are as follows:

- (1) Volatility from high frequency data can be aggregated, whereas aggregated data at low frequencies typically display no volatility;
- (2) Enables prediction of uncertainty regarding the imposition of tourism taxes on international tourist arrivals;
- (3) Enables prediction of uncertainty regarding the imposition of environmental taxes on polluters;
- (4) Enables prediction of risk in finance;
- (5) Enables the prediction of dynamic correlations for constructing financial portfolios;
- (6) Enables the computation of dynamic confidence intervals;
- (7) Enables the computation of dynamic Value-at-Risk (VaR) thresholds;
- (8) Enables the computation of dynamic confidence intervals for dynamic VaR thresholds;
- (9) Enables the prediction of dynamic variances and covariances for constructing dynamic VaR thresholds for financial portfolios;
- (10) Enables the derivation of a strategy for optimizing dynamic VaR.

6. The Ten Commandments for Optimizing Value-at-Risk and Daily Capital Charges

Credit risk is the most important type of risk in terms of monetary value, while market

risk is typically concerned with stocks and bonds, and related financial derivatives, as well as exchange rates and interest rates. Operational risk involves credit and market risk, as well as other operational aspects or risk.

For market, credit and operational risk, choose wisely from:

- (C1) Conditional, stochastic and realized volatility.**
- (C2) Symmetry, asymmetry and leverage.**
- (C3) Dynamic correlations and dynamic covariances.**
- (C4) Single index and portfolio models.**
- (C5) Parametric, semiparametric and nonparametric models.**
- (C6) Estimation, simulation and calibration of parameters.**
- (C7) Assumptions, regularity conditions and statistical properties.**
- (C8) Accuracy in calculating moments and forecasts.**
- (C9) Optimizing threshold violations and economic benefits.**
- (C10) Optimizing private and public benefits of risk management.**

The commandments progress from the theoretical to the practical. The remainder of this section briefly discusses each self-explanatory commandment.

(C1) Choose Wisely from Conditional, Stochastic and Realized Volatility

Different volatility models provide different estimates and forecasts of risk, while different data frequencies lead to different choices of volatility model.

- (i) Conditional volatility, which is typically used to model monthly and daily

data, is latent (see Li, Ling and McAleer (2002) and Bauwens, Laurent and Rombouts (2006) for recent reviews). The ease of computation, even for some multivariate models, as well as its availability in several widely used econometric software packages, has made this class of models very popular.

- (ii) Stochastic volatility is typically used to model and forecast daily data, and is also latent (see Asai, McAleer and Yu (2006) for a recent review of multivariate stochastic volatility models). A distinct advantage of stochastic volatility models is the incorporation of leverage, at least at the univariate level. However, the computational burden can be quite severe, especially for multivariate processes. Moreover, standard econometric software packages do not yet seem to have incorporated stochastic volatility algorithms.
- (iii) Realized volatility is observable, but is subject to measurement error (see McAleer and Medeiros (2008a) for a recent review). Such models are used to calculate observed volatility using tick data. The generated realized daily volatility measures are then typically modelled using a wide variety of long memory or fractionally integrated time series processes. It should be emphasized that microstructure noise (or measurement error) is a standard problem that arises where realized volatility is used as an estimate of daily integrated volatility.

(C2) Choose Wisely from Symmetry, Asymmetry and Leverage

Distinguish carefully between asymmetry and leverage, and models that incorporate asymmetry and leverage, either by construction or through the use of parametric

restrictions.

Asymmetry is a straightforward concept, but leverage seems to be the subject of much misunderstanding in practice.

Definition: Asymmetry captures the different impacts of positive and negative shocks of equal magnitude on volatility.

Definition: Leverage captures the effects of negative (positive) shocks of equal magnitude on increasing (decreasing) the debt-equity ratio, thereby increasing (decreasing) subsequent volatility and risk (see Black (1976)).

Thus, leverage is a special case of asymmetry, with volatility decreasing progressively as returns shocks change progressively from negative to positive.

It follows that symmetry is the absence of asymmetry, including leverage. The widely used ARCH and GARCH models are symmetric.

Several popular models of volatility display asymmetry, though not necessarily leverage. For example, the GJR model of Glosten, Jagannathan and Runkle (1992) is asymmetric but does not display leverage (see McAleer, Hoti and Chan (2009) for a multivariate extension of the asymmetric GJR model, VARMA-AGARCH), while the EGARCH model of Nelson (1991) is asymmetric and can display leverage, depending on the signs of the coefficients relating to the size and sign effects.

In the context of stochastic volatility models, leverage is imposed through a negative

correlation between returns and subsequent volatility shocks. There are many alternative types of asymmetry at the univariate and multivariate levels (see, for example, Asai and McAleer (2005, 2009b) and McAleer and Medeiros (2008b)).

Although realized volatility models may be modified to incorporate both intra- and inter-day leverage, this does not seem to have been accomplished to date.

Extensions of leverage to multivariate processes are somewhat more difficult.

(C3) Choose Wisely between Dynamic Correlations and Dynamic Covariances

Covariances and correlations are used to model relationships between pairs of assets for portfolio risk management:

Correlations are used in the construction of a portfolio. Consider two financial assets, X and Z:

- (a) correlation $(X,Z) \rightarrow +1 \rightarrow$ specialize on the asset with the higher historical and/or expected returns;
- (b) correlation $(X,Z) \rightarrow -1 \rightarrow$ hedge (or diversify);
- (c) correlation $(X,Z) \rightarrow 0 \rightarrow$ need to balance risk and returns.

On the other hand, dynamic variances and covariances are used to calculate the VaR of a given portfolio.

As a multivariate extension of equation (3), consider the relationship between

covariances and correlations, as follows:

$$Q_t = D_t \Gamma_t D_t, \quad (5)$$

in which Q_t is the conditional covariance matrix, Γ_t is the conditional correlation matrix, and D_t is the diagonal matrix of conditional standard deviations, namely $\sqrt{h_{it}}$, where

$$h_{it} = \omega_i + \alpha_i \varepsilon_{i,t-1}^2 + \beta_i h_{i,t-1},$$

and $i = 1, \dots, m$ is the number of assets in the portfolio.

The matrix Q_t is used to calculate VaR forecasts, while the matrix Γ_t is used to construct and update portfolios.

It should be emphasized that Q_t can be modelled directly, as in the BEKK model of Engle and Kroner (1995), or indirectly through modelling Γ_t , and using (5), such as using the dynamic conditional correlation (DCC) model of Engle (1992), of which the constant conditional correlation (CCC) model of Bollerslev (1990) is a special case (for an application of the scalar BEKK versus indirect DCC models, see Caporin and McAleer (2008)).

The BEKK model is given as follows:

$$Q_t = QQ' + A\varepsilon_{t-1}\varepsilon_{t-1}'A' + BQ_{t-1}B', \quad (6)$$

where Q_t , Q , A , B are m dimensional matrices. It is clear that the specification in (6) guarantees a positive definite covariance matrix. However, as the dimensions of the three parameter matrices, namely Q , A , B , are the same as for Q_t , a computational difficulty can arise frequently in the BEKK model of Q_t as it suffers from the so-called “curse of dimensionality” in having far too many parameters.

It follows from (5) that

$$\Gamma_t = D_t^{-1}Q_tD_t^{-1}, \quad (7)$$

so that Γ_t can be modelled directly, or indirectly through modelling Q_t . A parsimonious model of Γ_t is given in the DCC model, as follows:

$$\Gamma_t = (1 - \theta_1 - \theta_2) + \theta_1\eta_{i,t-1}\eta_{i,t-1}' + \theta_2\Gamma_{t-1}, \quad (8)$$

where θ_1, θ_2 are scalar parameters. As the specification does not guarantee that the elements along the main diagonal are all unity, and all of the off-diagonal terms lie in the range $[-1, 1]$, Engle (2002) standardizes the matrix in (8) so that the elements satisfy the definition of a conditional correlation.

Although the DCC model is parsimonious in terms of parameters, a common empirical finding, especially for stock indexes, is that the long run conditional correlation matrix is constant (namely, $\theta_1 = 0, \theta_2 = 1$), with the outcome being that news has little practical effect in changing the purportedly dynamic conditional correlations. This is not altogether surprising as, apart from the positive diagonal elements in a matrix version of DCC (see, for example, the GARCC model of McAleer, Chan, Hoti and Lieberman (2008)), the off-diagonal terms of the coefficient matrix can be positive or negative. The imposition of an unrealistic constraint that all the elements of the matrix are the same constant has the effect of making θ_1 very close to zero and θ_2 very close to unity, especially for a portfolio with a large numbers of assets. This practical problem associated with DCC is alleviated with various extensions of the the model, such as GARCC (see also Caporin and McAleer (2009a)).

Dynamic correlations have recently been developed for multivariate stochastic volatility models using the Wishart distribution (see Asai and McAleer (2009a)), but the computational burden can be quite severe.

To date there does not seem to have been any research undertaken on modelling the dynamic correlations for pairs of standardized residuals for multivariate realized volatility models.

(C4) Choose Wisely between Index and Portfolio Models

Estimation and forecasting of VaR thresholds of a portfolio requires estimation and forecasting the variance and covariances of portfolio returns. Volatility models can be

used to estimate the variance of portfolio returns either by (1) fitting a univariate volatility model to the portfolio returns (hereafter called the single index model (see McAleer and da Veiga (2008a, 2008b) and Asai and McAleer (2008)); or (2) using a multivariate volatility model to forecast the conditional variance of each asset in the portfolio to calculate the forecasted portfolio variance (hereafter called the portfolio model). Asai and McAleer (2009a) have extended the idea of a portfolio index model to a multivariate GARCH process.

The central issue is the signal to noise ratio. Single index models require only a univariate model to calculate the variance of the single index, and hence require no covariances or correlations. There is little signal, but there is also little noise. Portfolio models have a lot of signal because of all the pairs of covariances and correlations, but there is also a lot of noise in the estimated parameters or coefficients.

As an illustration in the case of two assets, let $y_t = \lambda x_t + (1 - \lambda)z_t$,

where

y_t = single index return to the portfolio of two assets,

x_t = return to financial asset, X_t

z_t = return to financial asset, Z_t

$\lambda x_t + (1 - \lambda)z_t$ = portfolio return on two assets,

$$h_{y_t} = \omega_y + \alpha_y \varepsilon_{y_{t-1}}^2 + \beta_y h_{y_{t-1}}$$

$$h_{x_t} = \omega_x + \alpha_x \varepsilon_{x_{t-1}}^2 + \beta_x h_{x_{t-1}}$$

$$h_{z_t} = \omega_z + \alpha_z \varepsilon_{z_{t-1}}^2 + \beta_z h_{z_{t-1}}$$

$$h_{xz_t} = \omega_{xz} + \alpha_{xz} \varepsilon_{x_{t-1}} \varepsilon_{z_{t-1}} + \beta_{xz} h_{xz_{t-1}}.$$

As the two conditional variances and single covariance are estimated, it follows that:

$$h_{y_t} \neq \lambda^2 h_{x_t} + (1 - \lambda)^2 h_{z_t} + 2\lambda(1 - \lambda) h_{xz_t},$$

even though

$$\text{var}(y_t) = \lambda^2 \text{var}(x_t) + (1 - \lambda)^2 \text{var}(z_t) + 2\lambda(1 - \lambda) \text{cov}(x_t, z_t).$$

Therefore, the single index and portfolio approaches can lead to different results when the time-varying variances and covariances are estimated.

The number of covariances increases dramatically with m , the number of assets in the portfolio. Thus, for $m = 2, 3, 4, 5, 10, 20, 30, 40, 50, 100$, the number of covariances is 1, 3, 6, 10, 45, 190, 435, 780, 1225 and 4950, respectively. This increases the computational burden significantly, unless some structure can be imposed to increase parsimony.

(C5) Choose Wisely from Parametric, Semiparametric and Nonparametric Models

Conditional volatility is latent, and conditional volatility models are typically semiparametric or parametric.

Stochastic volatility is latent, and stochastic volatility models are typically parametric.

Realized volatility is observable, and realized volatility models are nonparametric.

Each type of model typically has an optimal method of estimation, as follows:

(1) Conditional volatility models are typically estimated by the maximum likelihood (ML) method when the standardized residuals are normally distributed, or by the quasi-maximum likelihood (QML) method when they are not normal (for further details, see Li, Ling and McAleer (2002) or Bauwens, Laurent and Rombouts (2006));

(2) Stochastic volatility models are typically estimated by the Bayesian Markov Chain Monte Carlo (MCMC), Monte Carlo Likelihood (MCL), Empirical Likelihood (EL), or Efficient Method of Moments (EMM) methods (for further details, see Asai, McAleer and Yu (2006));

(3) Realized volatility models are typically estimated by nonparametric methods (for detailed analyses, see Zhang, Mykland and Aït-Sahalia (2005), Aït-Sahalia, Mykland and Zhang (2006), and Barndorff-Nielsen, Hansen, Lunde Shephard (2008); for recent reviews of the literature, see Bandi and Russell (2007) and McAleer and Medeiros (2008a)).

(C6) Choose Wisely from Estimation, Simulation and Calibration of Parameters

RiskmetricsTM calibrates the parameters in both univariate and multivariate models, and hence is concerned with forecasting volatility. As there are no estimated parameters, there are no standard errors, and there can be no inference.

Conditional volatility models are concerned with estimation and forecasting. The statistical properties of consistency and asymptotic normality for univariate and multivariate conditional volatility models are now well established in the literature (see, for example, Ling and McAleer (2002a, 2002b, 2003)).

Stochastic volatility models are concerned with simulation and forecasting. As the Bayesian MCMC, MCL, EL or EMM methods are typically used for estimation, the small sample or asymptotic properties of the estimators are well known.

Realized volatility models are based on nonparametric estimation methods, and their asymptotic properties are well known. Large samples are typically required for the use of nonparametric methods. In analyzing ultra high frequency tick data on financial returns, very large samples are usually available.

(C7) Choose Wisely from Assumptions, Regularity Conditions and Statistical Properties

It is essential to distinguish between assumptions and regularity conditions that are derived from the assumptions of the underlining model.

Assumptions are required to obtain the moment conditions (otherwise known as

regularity conditions), especially the second and fourth moments, as well as log-moments. Moment conditions are required to obtain the statistical properties of consistency and asymptotic normality, thereby also providing diagnostic checks of the model.

The regularity conditions for univariate and multivariate GARCH models are now well known in the literature. The log-moment conditions for GARCH(1,1) and GJR(1,1) were established by Lee and Hansen (1994) and McAleer, Chan and Marinova (2007), respectively, and the second and fourth moments for the multivariate extensions of GARCH and GJR, namely the VARMA-GARCH and VARMA-AGARCH models, by Ling and McAleer (2003) and McAleer, Hoti and Chan (2009)). The statistical properties of BEKK were established by Jeantheau (1998) and Comte and Lieberman (2003), and for a generalization of DCC, namely the Generalized Autoregressive Conditional Correlation (GARCC) model, by McAleer, Chan, Hoti and Lieberman (2008).

The Bayesian statistical properties of the univariate and multivariate stochastic volatility models are well known, as they are also for realized volatility models.

(C8) Choose Wisely between Accuracy in Calculating Moments and Forecasts

Derive appropriate measures to determine the accuracy of estimates and forecasts, including economic benefits, and evaluate dynamic checks of second moments and log-moment conditions for univariate and multivariate conditional volatility models.

The moment and log-moment conditions should be calculated as diagnostic checks of

the underlying volatility models. The second moment conditions for univariate GARCH and GJR are straightforward to calculate, but the log-moments are typically not calculated. The corresponding moment and log-moment conditions for multivariate processes are typically ignored in practice. Moment conditions are not necessary for the EGARCH model as it is a function of the standardized residuals rather than the unconditional returns shocks.

Typically, only the second moment conditions for univariate ARCH and GARCH models are checked, despite the ease of computing the log-moment conditions for GARCH(1,1) and GJR(1,1) (see Lee and Hansen (1994) and McAleer, Chan and Marinova (2007), respectively).

The accuracy of estimating models of the respective underlying processes should be weighed against any improvements in forecasting performance.

(C9) Choose Wisely between Optimizing Threshold Violations and Economic Benefits

As the daily capital charge is to be minimized with respect to both k and VaR, when the estimated volatility increases, the penalty from violation (k) tends to decrease, while VaR increases, with the end result being that capital charges increase.

Similarly, when the estimated volatility decreases, the penalty from violation (k) tends to increase, while VaR decreases, with the end result being that capital charges decrease.

Therefore, daily capital charges should be minimized, and economic benefits should be maximized, using k and VaR as an endogenous decision variables.

Tables 2 and 3 make it clear that increasing the number of violations leads to lower mean daily capital charges across all the volatility models considered, and for both the single index and portfolio approaches.

The Basel II penalty structure would seem to be too lenient for violations, especially for large violations, as the penalty structure focuses on the number, rather than the number and magnitude, of violations in (1).

A strategy to minimize daily capital charges has been devised by McAleer, Pérez-Amaral and Jiménez-Martin (2009), with k and VaR as endogenous choice variables. The strategy seems to work much better than treating k as exogenous. McAleer, Jiménez-Martin and Perez Amaral (2009) assess whether the Basel II Accord has encouraged risk management during the 2008-09 financial crisis, and evaluate alternative conservative and aggressive risk management strategies.

In short, there is a need to change the penalty structure under the Basel Accord, otherwise there is likely to be continuing excessively high risk taking.

(C10) Choose Wisely between Optimizing Private and Public Benefits of Risk Management

In summary, ADIs should balance violations (which may be crucial to avoid for public relations purposes) against daily capital charges.

As the regulations stand at present, the Basel II Accord would seem to encourage risk taking as the penalties for violations are exceedingly low, and hence would seem to favour violations rather than managing and monitoring excessive risk taking (see Tables 2 and 3).

It is essential to determine whether increasing the number of violations, which subsequently leads to lower mean daily capital charges, is in the public interest in terms of monitoring and managing excessive risk and VaR.

There is an urgent requirement to balance risk taking with prudent self regulation or government regulation in the banking and finance industry (as seen earlier in the 2003 remarks by Alan Greenspan). There may, in fact, be moral hazard in self regulation in the banking and finance industry, as qualified by Greenspan in 2003, especially in terms of federal guarantees of private debt.

The financial meltdown in September and October 2008 (and still counting) demands more careful and responsible regulation of the industry.

7. Concluding Remarks

The largely self-regulated management of excessive risk taking in the world of banking and finance has led to the worst financial disaster in September-October 2008 since the market collapse of 1929. Government and private regulation of credit risk, which is the most important type of risk in terms of monetary value, and market risk, which is typically concerned with stocks and bonds, related financial derivatives, and

exchange rates and interest rates, have been found to be largely inadequate.

This paper analysed market risk management and monitoring under the Basel II Accord, and presented the optimization problem facing authorized deposit-taking institutions. Ten reasons for modelling time-varying variances, covariances and correlations using high and ultra high frequency data were given, and the Ten Commandments for optimizing Value-at-Risk (VaR) and daily capital charges were analysed. As the Basel II Accord would seem to encourage risk taking at the expense of providing accurate measures and forecasts of risk and VaR, the paper gave some idiosyncratic suggestions and guidelines regarding how improvements might be made regarding optimal strategies for risk monitoring and management, especially when there is a strong and understandable movement into holding and managing cash than in dealing with risky financial investments.

8. Epilogue

“Rules are made to be broken.”

Anonymous

“Ignore all rules.”

Anonymous

“One of the quickest ways to break something is to fix it when it ain’t broken.”

Anonymous

Butch Cassidy: “Every day you get older. Now that’s a law.”

Butch Cassidy and the Sundance Kid

Etta Place: “Do you know what you’re doing?”

Butch Cassidy: “Theoretically.”

Butch Cassidy and the Sundance Kid

Commandments, laws, rules, regulations, guidelines, codes, call them what you will. They are all made to be ignored and/or broken. The frequency of breaking commandments may be testament to the indefatigability of human beings to ignore the blindingly obvious, or to the blindingly obvious fact that such commandments may not always be particularly useful.

The fact remains that commandments are routinely ignored, for whatever reason. The Ten Commandments for organizing a conference (McAleer (1997)), for attending a conference (McAleer and Oxley (2001)), for presenting a conference paper (McAleer and Oxley (2002)), for ranking university quality (McAleer (2005b)), for academics (McAleer and Oxley (2005)), for managing investments (Caporin and McAleer (2009)), and for managing value-at-risk under the Basel II Accord (Jiménez-Martin et al. (2009)), are frequently downloaded (using download statistics) and read, but are also frequently ignored, although the financial penalties are not quite as frightening as when the Ten Commandments for optimizing VaR and daily capital charges are broken.

Table 2: Mean Daily Capital Charge and AD of Violations for the Single Index Model

Model	Number of Violations	Mean Daily Capital Charge	AD of Violations	
			Maximum	Mean
Standardized Normal	35	12.329	3.506	0.631
Riskmetrics™	27	9.113	2.772	0.456
ARCH	62	8.319	2.758	0.552
ARCH-t	36	8.832	2.302	0.551
GARCH	34	8.095	2.430	0.464
GARCH-t	14	8.981	2.302	0.441
GJR	34	8.095	2.430	0.464
GJR-t	13	9.903	1.701	0.521
PGARCH	31	8.041	1.205	0.362
PGARCH-t	9	9.034	1.708	0.510
EGARCH	30	7.968	1.154	0.298
EGARCH-t	9	8.986	1.556	0.489
Notes:				
(1) The daily capital charge is given as the negative of $(3+k)$ times the greater of the previous day's VaR or the average VaR over the last 60 business days, where k is the violation penalty.				
(2) AD is the absolute deviation of the violations from the VaR forecast.				

Note: This is Table 8 in McAleer and da Veiga (2008b).

The models are as follows: ARCH was developed by Engle (1982), GARCH by Bollerslev (1986), GJR by Glosten, Jagannathan and Runkle (1992), EGARCH by Nelson (1991), and PGARCH is the asymmetric power GARCH model of Ding, Granger and Engle (1993).

Table 3: Mean Daily Capital Charge and AD of Violations for the Portfolio Model

Model	Number of Violations	Mean Daily Capital Charge	AD of Violations	
			Maximum	Mean
Standardized Normal	36	12.916	3.509	0.617
Riskmetrics TM	28	8.509	2.516	0.413
ARCH	19	9.132	2.271	0.540
ARCH-t	9	10.581	1.691	0.463
CCC	7	9.685	2.125	0.498
CCC-t	1	11.498	1.489	1.489
GJR	7	9.724	1.657	0.505
GJR-t	2	11.571	0.857	0.549
EGARCH	6	9.692	1.566	0.466
EGARCH-t	2	11.544	0.727	0.482
PGARCH	6	9.787	1.485	0.472
PGARCH-t	2	11.658	0.623	0.490
VARMA-GARCH	6	9.760	1.974	0.454
VARMA-GARCH-t	1	11.633	1.287	1.287
PS-GARCH	7	10.700	1.902	0.442
PS-GARCH-t	1	11.833	1.321	1.321
Notes:				
(1) The daily capital charge is given as the negative of $(3+k)$ times the greater of the previous day's VaR or the average VaR over the last 60 business days, where k is the violation penalty.				
(2) AD is the absolute deviation of the violations from the VaR forecast.				

Note: This is Table 9 in McAleer and da Veiga (2008b).

In addition to the models described in the note to Table 2, the models are as follows: CCC was developed by Bollerslev (1990), VARMA-GARCH by Ling and McAleer (2003), and PS-GARCH by McAleer and da Veiga (2008a).

References

- Aït-Sahalia, Y., P.A. Mykland and L. Zhang (2006), Ultra high Frequency Volatility Estimation with Dependent Microstructure Noise, Working Paper, w11380, NBER.
- Asai, M. and M. McAleer (2005), Dynamic Asymmetric Leverage in Stochastic Volatility Models, *Econometric Reviews*, 24, 317-332.
- Asai, M. and M. McAleer (2008), A Portfolio Index GARCH Model, *International Journal of Forecasting*, 24, 449-461.
- Asai, M. and M. McAleer (2009a), The Structure of Dynamic Correlations in Multivariate Stochastic Volatility Models, *Journal of Econometrics*, 150, 182-192.
- Asai, M. and M. McAleer (2009b), Alternative Asymmetric Stochastic Volatility Models, to appear in *Econometric Reviews*.
- Asai, M., M. McAleer and J. Yu (2006), Multivariate Stochastic Volatility: A Review, *Econometric Reviews*, 25, 145-175.
- Bandi, F.M. and J.R. Russell (2007), Volatility Estimation, in J.R. Birge and V. Linetsky (eds.), *Handbook of Financial Engineering*, Elsevier, Amsterdam.
- Barndorff-Nielsen, O.E., P.H. Hansen, A. Lunde and N. Shephard (2008), Subsampling realised kernels, to appear in *Econometrica*.
- Basel Committee on Banking Supervision, (1988), "International Convergence of Capital Measurement and Capital Standards", BIS, Basel, Switzerland.
- Basel Committee on Banking Supervision, (1995), "An Internal Model-Based Approach to Market Risk Capital Requirements", BIS, Basel, Switzerland.
- Basel Committee on Banking Supervision, (1996), "Supervisory Framework for the Use of "Backtesting" in Conjunction with the Internal Model-Based Approach to Market Risk Capital Requirements", BIS, Basel, Switzerland.

- Bauwens, L., S. Laurent and V.K. Rombouts (2006), Multivariate GARCH: A Survey, *Journal of Applied Econometrics*, 21, 79-109.
- Black, F. (1976), Studies of Stock Market Volatility Changes, in 1976 Proceedings of the American Statistical Association, Business and Economic Statistics Section, pp. 177-181.
- Bollerslev, T. (1986), Generalised Autoregressive Conditional Heteroscedasticity, *Journal of Econometrics*, 31, 307-327.
- Bollerslev, T. (1990), Modelling the Coherence in Short-run Nominal Exchange Rates: A Multivariate Generalized ARCH Approach, *Review of Economics and Statistics*, 72, 498-505.
- Caporin, M. and M. McAleer (2008), Scalar BEKK and Indirect DCC, *Journal of Forecasting*, 27, 537-549.
- Caporin, M. and M. McAleer (2009a), Do We Really Need Both BEKK and DCC? A Tale of Two Covariance Models, Available at SSRN: <http://ssrn.com/abstract=1338190>.
- Caporin, M. and M. McAleer (2009b), The Ten Commandments for Managing Investments, Available at SSRN: <http://ssrn.com/abstract=1342265>.
- Comte, F. and O. Lieberman (2003), Asymptotic Theory for Multivariate GARCH Processes, *Journal of Multivariate Analysis*, 84, 61-84.
- Ding, Z., C.W.J. Granger and R.F. Engle (1993), A Long Memory Property of Stock Market Returns and a New Model, *Journal of Empirical Finance*, 1, 83-106.
- Engle, R.F. (1982), Autoregressive Conditional Heteroscedasticity with Estimates of the Variance of United Kingdom Inflation, *Econometrica*, 50, 987-1007.
- Engle, R.F. (2002), Dynamic Conditional Correlation: A Simple Class of Multivariate Generalized Autoregressive Conditional Heteroskedasticity Models, *Journal of Business and Economic Statistics*, 20, 339-350.
- Engle, R.F. and K.F. Kroner (1995), Multivariate Simultaneous Generalized ARCH,

- Econometric Theory, 11, 122-150.
- Engle, R.F. and S. Manganelli (2002), .CAViaR: Conditional Autoregressive Value at Risk by Regression Quantiles, *Journal of Business and Economic Statistics*, 22, 367-381.
- Glosten, L.R., R. Jagannathan and D.E. Runkle (1992), On the Relation Between the Expected Value and Volatility of the Nominal Excess Return on Stocks, *Journal of Finance*, 46, 1779-1801.
- Jeantheau, T. (1998), Strong Consistency of Estimators for Multivariate ARCH Models, *Econometric Theory*, 14, 70-86.
- Jiménez-Martin, J.-A., M. McAleer and T. Perez Amaral (2009), The Ten Commandments for Managing Value-at-Risk Under the Basel II Accord, Available at SSRN: <http://ssrn.com/abstract=1356803>.
- Jorion, P. (2000), *Value at Risk: The New Benchmark for Managing Financial Risk*, McGraw-Hill, New York.
- Lee, S.W. and B.E. Hansen (1994), Asymptotic Theory for the GARCH(1,1) Quasi-Maximum Likelihood Estimator, *Econometric Theory*, 10, 29-52.
- Li, W.K., S. Ling and M. McAleer (2002), Recent Theoretical Results for Time Series Models with GARCH Errors, *Journal of Economic Surveys*, 16, 245-269. Reprinted in M. McAleer and L. Oxley (eds.), *Contributions to Financial Econometrics: Theoretical and Practical Issues*, Blackwell, Oxford, 2002, pp. 9-33.
- Ling, S. and M. McAleer (2002a), Necessary and Sufficient Moment Conditions for the GARCH(r,s) and Asymmetric Power GARCH(r,s) Models, *Econometric Theory*, 18, 722-729.
- Ling, S. and M. McAleer (2002b), Stationarity and the Existence of Moments of a Family of GARCH Processes, *Journal of Econometrics*, 106, 109-117.
- Ling, S. and M. McAleer (2003), Asymptotic Theory for a Vector ARMA-GARCH Model, *Econometric Theory*, 19, 278-308.

- McAleer, M. (1997), The Ten Commandments for Organizing a Conference, *Journal of Economic Surveys*, 11, 231-233.
- McAleer, M. (2005a), Automated Inference and Learning in Modeling Financial Volatility, *Econometric Theory*, 21, 232-261.
- McAleer, M. (2005b), The Ten Commandments for Ranking University Quality, *Journal of Economic Surveys*, 19, 649-653.
- McAleer, M., F. Chan, S. Hoti and O. Lieberman (2008), Generalized Autoregressive Conditional Correlation, *Econometric Theory*, 24, 1554-1583.
- McAleer, M., F. Chan and D. Marinova (2007), An Econometric Analysis of Asymmetric Volatility: Theory and Application to Patents, *Journal of Econometrics*, 139, 259-284.
- McAleer, M., S. Hoti and F. Chan (2009), Structure and Asymptotic Theory for Multivariate Asymmetric Conditional Volatility, *Econometric Reviews*, 28, 422-440.
- McAleer, M., J.-A. Jiménez-Martin and T. Perez Amaral (2009), Has the Basel II Accord Encouraged Risk Management During the 2008-09 Financial Crisis?, Available at SSRN: <http://ssrn.com/abstract=1397239>.
- McAleer, M., T. Pérez-Amaral and J.-Á. Jiménez-Martin (2009), A Decision Rule to Minimize Daily Capital Charges in Forecasting Value-at-Risk, Available at SSRN: <http://ssrn.com/abstract=1349844>.
- McAleer, M. and M.C. Medeiros (2008a), Realized Volatility: A Review, *Econometric Reviews*, 27, 10-45.
- McAleer, M. and M.C. Medeiros (2008b), A Multiple Regime Smooth Transition Heterogeneous Autoregressive Model for Long Memory and Asymmetries, *Journal of Econometrics*, 147, 104-119.
- McAleer, M. and L. Oxley (2001), The Ten Commandments for Attending a Conference, *Journal of Economic Surveys*, 15, 671-678.

- McAleer, M. and L. Oxley (2002), The Ten Commandments for Presenting a Conference Paper, *Journal of Economic Surveys*, 16, 215-218.
- McAleer, M. and L. Oxley (2005), The Ten Commandments for Academics, *Journal of Economic Surveys*, 19, 823-826.
- McAleer, M. and B. da Veiga (2008a), Forecasting Value-at-Risk with a Parsimonious Portfolio Spillover GARCH (PS-GARCH) Model, *Journal of Forecasting*, 27, 1-19.
- McAleer, M. and B. da Veiga (2008b), Single Index and Portfolio Models for Forecasting Value-at-Risk Thresholds, *Journal of Forecasting*, 27, 217-235.
- Nelson, D. (1991), Conditional Heteroskedasticity in Asset Returns: A New Approach, *Econometrica*, 59, 347-370.
- Pesaran, M.H., C. Schleicher and P. Zaffaroni (2008), Model Averaging in Risk Management with an Application to Futures Markets, to appear in *Journal of Empirical Finance*.
- Riskmetrics™ (1996), J.P. Morgan Technical Document, 4th edition, New York. J.P. Morgan.
- Zhang, L., P.A. Mykland and Y. Aït-Sahalia (2005), A Tale of Two Time Scales: Determining Integrated Volatility with Noisy High Frequency Data, *Journal of the American Statistical Association*, 100, 1394 – 1411.